Efficient CI Estimation for Neutrino Oscillation Parameters with Gaussian Process

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Neutrino Oscillations

- Neutrino oscillations between flavor states occur with a well defined probability which depends on the U_{PMNS} mixing matrix
- ▶ LBL experiments (focus of this talk) measure $P(\nu_{\mu} \rightarrow \nu_{\mu})$ and $P(\nu_{\mu} \rightarrow \nu_{e})$ to infer :
- $\Delta m_{32}^2 > 0$ or < 0? (Normal or Inverted)
 - Identifying mass hierarchy (NH or IH) has implications for neutrino mass measurements
- Octant of θ_{23} or $\theta_{23} = 45^{\circ}$?
- $sin\delta_{CP} \neq 0$?
 - Lepton sector CP-violation. Gives us a clue towards explaining matter-antimatter asymmetry



Statistical Issues

- Experiments collect only a handful of statistics. $\mathcal{O}(10-100)$ over years of operation for the $\nu_{\mu} \rightarrow \nu_{e}$ channel
- Complicated interplay between different parameters ⇒ difficult to delineate
- Confidence Intervals are hard to find as Likelihood ratios don't satisfy asymptotic properties.
- Let's illustrate this with a toy experiment..



Toy Experiment



- Create a toy NOvA-like experiment. Data (x) generated from Poisson variations at some chosen oscillation parameters.
- With (θ, δ) denoting list of oscillation and nuisance (flux and xsec errors) parameters,
- ▶ Best-fit $(\hat{\theta}, \hat{\delta})$ found by minimizing negative log-likelihood over energy bins, *i*

$$-2\log L(\theta,\delta) = -2\sum_{i\in I}\log \operatorname{Pois}(x_i; v(\theta,\delta)_i) - \sum_{i\in I}x_i + \sum_{i\in I}v(\theta,\delta)_i + \delta^2$$

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Confidence Intervals

- Likelihood Ratio Tests (LRT) (Δχ² from global best fit) typically used for estimating confidence intervals.
- ► In asymptotic case, test statistic : $\Delta \chi^2 \sim \chi_k^2 \implies$ look up significance from PDG (Wilks' Theorem)
- ► In others ⇒ Feldman-Cousins, i.e

- \blacktriangleright Explicitly simulate $\Delta\chi^2$ distribution using lots of pseudo-experiments
- Find p-value associated with $\Delta \chi^2_{data}$
- ▶ Gather all parameter values for which, say, percentile = 1 - p < 0.68 to get $1 - \sigma$ interval
- Correct coverage by construction
- Very heavy computational burden, often millions of CPU-hours needed!



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A more efficient FC

- In practice, FC proceeds via a grid search, for eg, simulating $\Delta \chi^2$ distributions for every point in $\sin^2 \theta_{23} \delta_{CP}$ space to find the 1- σ contour
- > If you had perfect foresight however, only the 1- σ boundaries are needed, but obviously not known apriori



- Can we get an idea of how this surface looks like with a few pseudo-experiment throws?
- Can we then use this approximate surface to tell us where those boundaries lie?

Gaussian Process

- Special case of Bayesian Learning. Assume p-value approximation is a random variable with a multivariate gaussian distribution
- ▶ We say, $f \sim \mathcal{GP}(\mu, k(\cdot, \cdot))$ if

$$\begin{pmatrix} f(x) \\ f(x') \end{pmatrix} \sim \mathcal{N}(\begin{bmatrix} \mu(x) \\ \mu(x') \end{bmatrix}, \begin{bmatrix} k(x,x) & k(x,x') \\ k(x,x') & k(x',x') \end{bmatrix}).$$

- Intuitively, we can picture each draw from a GP(µ, k(·, ·)) giving us a different f(x) with the average result being µ(x)
- ► The kernel encodes the correlation between nearby points. A commonly used kernel is the radial basis function, $k(x, x') = \exp(-(x x')^2/l^2)$
- ▶ A RBF kernel tells us that *GP* results at nearby points are highly influenced by observations at a given point while further out, they aren't.

Why $\mathcal{GP}s$?

- Enormously flexible! Can basically approximate any well behaved function with an appropriate choice of the kernel.
- Predictions at new data points are posterior distributions calculated with basic linear algebra, i.e for GP(0, k(·, ·)) :

$$f(x')|f(x) \sim \mathcal{N}(\frac{k(x,x')}{k(x,x)}f(x), k(x',x') - \frac{k(x,x')^2}{k(x,x)})$$



Kernel hyperparameters can be learned and updated iteratively as well

Optimised Confidence Interval Search

Use a priority score that guides the CI search in θ-space based on GP approximated p-value surface.

$$\mathsf{a}(heta) = \sum_{lpha_i} |rac{\sigma_{\hat{q}(heta)}}{\hat{q}(heta) - lpha_i}|$$

- Here, $\hat{q}(\theta)$ is \mathcal{GP} mean, $\sigma_{\hat{q}(\theta)}$ is \mathcal{GP} uncertainty, α_i is chosen to be (0.68, 0.90)
- $a(\theta)$ balances between exploration, i.e MC experiments at new points and exploitation, i.e reducing \mathcal{GP} error

8 / 28





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9 / 28





- ▶ "Real" data similar to latest best-fit estimate from NOvA. ($sin^2\theta_{23} = 0.56$, $\Delta m_{32}^2 = 2.44 \times 10^{-3} eV^2$, $\delta_{CP} = 1.5\pi$)
- $sin^2\theta_{23} \delta_{CP}$ 68% and 90% CI for IH after 5 iterations



- Grayscale denotes number of experiments thrown in relation to FC (2000)
- Algorithm does a good job of finding the FC contour edge!

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- $sin^2\theta_{23} \delta_{CP}$ 68% and 90% CI for NH after 5 iterations



- > 200 different runs for "real" data at the same point as before.
- Use classification accuracy of all grid points, taking FC result as truth, to evaluate performance.
- \blacktriangleright Progress shows the search algorithm converges to the FC value \sim 10× faster for 2D case and \sim 5× for 1D case



- Median Accuracies for 1D is 100%, for 2D is > 99.5% (both NH, IH)
- \blacktriangleright Mean Accuracies for 1D is 98.5% (99.8%) for NH (IH), for 2D is > 99% (both NH, IH)

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Summary and Conclusions

- Neutrino oscillation experiments provide interesting test case for estimating frequentist confidence intervals
- LBL experiments typically proceed via Feldman-Cousins
- However, simulating Δχ² distributions across multi-dimensional parameter space requires huge computational resources
- ▶ We've studied a Bayesian approach using Gaussian processes on a toy LBL set-up
- Helps us estimate frequentist contour edges to quite a high accuracy without having to sample the entire parameter space!
- Order of magnitude gain in computation!
- See PRD publication for more details : Phys.Rev.D 101 (2020) 1, 012001
- All code with illustrative notebooks here : https://github.com/nitish-nayak/ToyNuOscCI, maintained by Lingge (linggeli7@gmail.com) and myself (nayakb@uci.edu)

Backup

Toy Experiment

- ▶ Modelled on NOvA. Baseline, L = 810km with ν_{μ} flux peaking at 2GeV
- ▶ $u_{\mu} \rightarrow \nu_{e}$ by multiplying toy shapes for flux, cross-section and oscillation probability.
- ▶ 10% normalisation errors on flux and xsec model



▶ $P(\nu_{\mu} \rightarrow \nu_{e})$ using 3-flavor PMNS with MSW corrections added for matter propagation.

- Similar setup for $u_{\mu} \rightarrow \nu_{\mu}$ to constrain $sin^2(2\theta_{23})$ and $|\Delta m^2_{32}|$ but with 2-flavor approximation
- $P(\nu_{\mu} \rightarrow \nu_{\mu}) \sim 1 sin^2(2\theta_{23})sin^2(\Delta m_{32}^2 L/4E)$

$\mathcal{GP}s$ for FC

- Fitting a GP to target p-value surface for a given contour. (Stochasticity of the target surface)
- "Observation" at a given point in parameter space, θ means simulating the LRT distribution and finding the p-value of $crit(\theta)$
- Choose a RBF Kernel with an additional term incorporating variance of p-value estimate at θ .



$$k(\cdot, \cdot) = k_{RBF}(\cdot, \cdot) + \sigma_p^2 I$$

- The additional variance encodes the binomial error resulting from throwing finite number of experiments to simulate the LRT distribution at θ
- Allows us to incorporate varying number of experiments thrown into the CI search, reducing computational burden further.

- ▶ "Real" data similar to latest best-fit estimate from NOvA. ($sin^2\theta_{23} = 0.56$, $\Delta m_{32}^2 = 2.44 \times 10^{-3} eV^2$, $\delta_{CP} = 1.5\pi$)
- Significance of rejecting δ_{CP} only after 5 iterations. (p-value converted to Z-score significance)



- Rasmussen and Williams has a good discussion about convergence to true functions in regression settings (typically using squared loss functions) : http://www.gaussianprocess.org/gpml/chapters/RW7.pdf
- \blacktriangleright Well behaved \implies expressible as a generalised fourier series of kernel eigenfunctions
- ▶ If kernel is non-degenerate, approximation is guaranteed to converge to true function
- If degenerate, convergence towards an L_2 approximation of the true function
- Rates of convergence typically depends on mean and kernel smoothness as well as smoothness of the true function

\mathcal{GP} Fitting

• Hyperparameters (\mathbf{w}) learned via maximising log marginal likelihood :

$$p(\mathbf{y}|\mathbf{X},\mathbf{w}) = \int p(\mathbf{y}|\mathbf{X},\mathbf{w},\mathbf{f})p(\mathbf{f}|\mathbf{X},\mathbf{w})d\mathbf{f}$$

Clearly,

$$\mathbf{f}|\mathbf{X},\mathbf{w}\sim\mathcal{N}(\mathbf{0},\mathcal{K}(\mathbf{X},\mathbf{w}))$$

Some algebra gives us :

$$-2\log p(\mathbf{y}|\mathbf{X},\mathbf{w}) = \mathbf{y}^{\mathsf{T}} \boldsymbol{K}^{-1} \mathbf{y} + \log |\boldsymbol{K}| + n \log 2\pi$$

- Minimising above equation gives us a good choice for w
- log|K| acts as a penalty term for complexity and therefore reduces overfitting to data

\mathcal{GP} for FC

- "Gaussian" not a statement of the underlying distribution of the test statistic, which can still be heavily non-Gaussian
- Rather, "Gaussianity" for a stochastic process generating the test statistic distributions. Stochasticity mostly from finite FC grid resolution or finite number of MC experiments for simulating the test statistic distribution
- ► Assumption we're making for this stochasticity is that it can be parameterised by a kernel describing the relationship between the distributions at neighbouring points ⇒ multi-variate gaussian
- Also important to note, no real statement about FC coverage or handling of nuisance parameters. Assumes FC gives desired level of coverage
- Confidence Intervals still with frequentist interpretation
- Bayesian interpretation for "classification probability" of points in parameter space for desired confidence regions
- A good summary would be "Accelerating Frequentist CI search by estimating CI edges through Bayesian ML"

- ▶ GPs in HEP : arXiv:1709.05681, M. Frate, K. Cranmer et al. Using GPs to describe background spectra in dijet resonance searches at the LHC non-parametrically.
- > Used in Astrophysics for modelling stochasticity of light yields in stars, active galactic nuclei etc
- Many other fields!

Pseudo-code

```
Algorithm 1 \mathcal{GP} iterative confidence contour finding
```

```
for each iteration t = 1, 2, \dots do
Propose new points in parameter space arg max<sub>\theta</sub> a(\theta)
for each point \theta' do
    Simulate likelihood ratio distribution
    for k = 1, 2, ... do
         Perform a pseudo experiment
         Maximize the likelihood with respect to (\theta, \delta)
         Maximize the likelihood with constraint \theta = \theta'
    end for
    Obtain critical value c(\theta')
end for
Update \mathcal{GP} approximation \hat{c}(\theta)
Update confidence contours
```

end for

Results : NH, $sin^2\theta_{23} - \delta_{CP}$





NH, $sin^2\theta_{23} - \delta_{CP}$



NH, $sin^2\theta_{23} - \delta_{CP}$



NH, $sin^2\theta_{23} - \Delta m_{32}^2$



NH, $sin^2\theta_{23} - \Delta m_{32}^2$



NH, $sin^2\theta_{23} - \Delta m_{32}^2$

